

# THE WIDTH OF GALTON-WATSON TREES CONDITIONED BY THE SIZE

MICHAEL DRMOTA\* AND BERNHARD GITTEBERGER\*

ABSTRACT. It is proved that the moments of the width of Galton-Watson trees of size  $n$  and with offspring variance  $\sigma^2$  are asymptotically given by  $(\sigma\sqrt{n})^p m_p$  where  $m_p$  are the moments of the maximum of the local time of a standard scaled Brownian excursion. This is done by combining a weak limit theorem and a tightness estimate. The method is quite general and we state some further applications.

## 1. INTRODUCTION

In this paper we are considering rooted trees which are family trees of a Galton-Watson branching process conditioned to have total progeny  $n$ . These trees are also called simply generated trees (see [35]). Without loss of generality we may assume that the offspring distribution  $\xi$  is given by

$$\mathbf{P}\{\xi = k\} = \frac{\tau^k \varphi_k}{\varphi(\tau)}, \quad (1)$$

where  $(\varphi_k; k \geq 0)$  is a sequence of non-negative numbers such that  $\varphi(t) = \sum_{k \geq 0} \varphi_k t^k$  has a positive or infinite radius of convergence  $R$  and  $\tau$  is an arbitrary positive number within the circle of convergence of  $\varphi(t)$ . These conditions in particular imply that all moments of  $\xi$  exist and that  $\tau < R$ . Due to conditioning on the total progeny and finiteness of moments it is no restriction if we confine ourselves to studying only the critical case, that is,  $\mathbf{E}\xi = 1$  which equivalently means that  $\tau$  satisfies  $\tau\varphi'(\tau) = \varphi(\tau)$ . The variance of  $\xi$  can also be expressed in terms of  $\varphi(t)$  and is given by

$$\sigma^2 = \frac{\tau^2 \varphi''(\tau)}{\varphi(\tau)}. \quad (2)$$

Note that the offspring distribution (1) can be interpreted as assigning weights to all trees defined by

$$\omega(T) = \prod_{k \geq 0} \varphi_k^{n_k(T)}$$

for a tree  $T$  having  $n$  nodes,  $n_k$  of which have out-degree  $k$ ,  $k \geq 0$ . Denote by  $|T|$  the number of nodes of such a tree and let  $a_n$  be the (weighted) number of all trees with  $n$  nodes, i.e.

$$a_n = \sum_{T: |T|=n} \omega(T).$$

---

*Date:* May 8, 2004.

\* Department of Discrete Mathematics and Geometry, Wiedner Hauptstr. 8-10/104, A-1040 Wien, Austria.

Then the corresponding generating function  $a(z) = \sum_{n \geq 0} a_n z^n$  satisfies the functional equation

$$a(z) = z\varphi(a(z)). \quad (3)$$

Denote by  $(L_n(t), t \geq 0)$  the sequence of the generation sizes of a Galton-Watson tree the total progeny of which is  $n$ . For non-integer  $t$  we define  $L_n(t)$  by linear interpolation:

$$L_n(t) = (\lfloor t \rfloor + 1 - t)L_n(\lfloor t \rfloor) + (t - \lfloor t \rfloor)L_n(\lfloor t \rfloor + 1), \quad t \geq 0.$$

We are interested in the width of such a tree which is defined by

$$w_n = \max_{t \geq 0} L_n(t).$$

This quantity attracted the interest of many authors. First, Odlyzko and Wilf [37] became interested in this tree parameter when studying the bandwidth

$$\beta(T) = \min_f \left( \max_{(u,v) \in E(T)} |f(u) - f(v)| \right)$$

of a tree  $T$ , where  $f$  is an assignment of distinct integers to the vertices of the tree. They showed for a tree with  $n$  vertices and height  $h(T)$  and width  $w(T)$  that

$$\frac{n-1}{2h(T)} \leq \beta(T) \leq 2w(T) - 1$$

and furthermore they showed that there exist positive constants  $c_1$  and  $c_2$  such that the estimate

$$c_1 \sqrt{n} < \mathbf{E} w_n < c_2 \sqrt{n \log n} \quad (4)$$

holds. The exact order of magnitude was left as an open problem. Aldous conjectured [1, Conj. 4] that  $L_n$  (suitably normalized) converges to Brownian excursion local time. This was first proved in [15], later by different methods by Kersting [29] and Pitman [38]. More precisely, set

$$l_n(t) = \frac{2}{\sigma \sqrt{n}} L_n \left( \frac{2t}{\sigma} \sqrt{n} \right)$$

and

$$l(t) = \lim_{\varepsilon \rightarrow 0} \frac{1}{\varepsilon} \int_0^1 I_{[t, t+\varepsilon]}(W(s)) ds,$$

where  $(W(s), 0 \leq s \leq 1)$  is the standard scaled Brownian excursion, then the above described limit theorem reads as follows:

**Theorem 1** ([15]). *Let  $\varphi(t)$  be the GF of a family of random trees. Assume that  $\varphi(t)$  has a positive or infinite radius of convergence  $R$ . Furthermore suppose that the equation  $t\varphi'(t) = \varphi(t)$  has a minimal positive solution  $\tau < R$ . Then we have*

$$(l_n(t), t \geq 0) \xrightarrow{w} (l(t), t \geq 0)$$

in  $C[0, \infty)$ , as  $n \rightarrow \infty$ .

Partial results go back to [9, 22, 27, 34, 41]. A density representation was computed in [25].

This result implies that  $w_n/\sigma\sqrt{n}$  weakly converges to the maximum of Brownian excursion local time, which was proved directly by Takács [40].

**Theorem 2** ([15]). *Under the assumptions of Theorem 1 we have*

$$\sup_{t \geq 0} l_n(t) \xrightarrow{w} \sup_{t \geq 0} l(t).$$

Thus this suggests (but does not imply)  $\sqrt{n}$  as correct order of magnitude in (4).

Note that the maximum of local time is well studied (cf. [28, 8, 18, 3, 34]). We have  $\sup_{t \geq 0} l(t) \stackrel{d}{=} 2 \sup_{0 \leq t \leq 1} W(t)$ , moreover it is theta-distributed, i.e.,

$$\mathbf{P} \left\{ \sup_{0 \leq t \leq 1} l(t) \leq x \right\} = 1 - 2 \sum_{k \geq 1} (x^2 k^2 - 1) e^{-x^2 k^2 / 2},$$

and

$$\mathbf{E} \left[ \left( \sup_{t \geq 0} l(t) \right)^p \right] = 2^{p/2} p(p-1) \Gamma\left(\frac{p}{2}\right) \zeta(p).$$

The purpose of this paper is to show that we have convergence of moments for  $\sup_{t \geq 0} l_n(t)$ , too. We formulate it in terms of the width  $w_n = \max_{t \geq 0} L_n(t) = (\sigma/2) \sqrt{n} \sup_{t \geq 0} l_n(t)$ .

**Theorem 3.** *Suppose that there exists a minimal positive solution  $\tau < R$  of  $t\varphi'(t) = \varphi(t)$ . Then the width  $w_n$  satisfies*

$$\mathbf{E} (w_n^p) = \sigma^p 2^{-p/2} p(p-1) \Gamma\left(\frac{p}{2}\right) \zeta(p) \cdot n^{p/2} \cdot (1 + o(1))$$

as  $n \rightarrow \infty$ .

It should be further mentioned that Chassaing and Marckert [6] used the relation of parking functions and rooted trees as well as the strong convergence theorem of Komlos, Major and Tusnady [33] to derive tight bounds for the moments of the width for Cayley trees. They showed (here and throughout the whole paper,  $a \ll b$  denotes  $a \leq Cb$  for some positive constant  $C$ )

**Theorem 4** ([6]). *If  $\varphi(t) = e^t$  and  $p \geq 1$ , then*

$$\left| \mathbf{E} \left( \frac{w_n}{\sigma \sqrt{n}} \right)^p - \mathbf{E} \left( \frac{1}{2} \sup_{t \geq 0} l(t) \right)^p \right| = \left| \mathbf{E} \left( \frac{w_n}{\sigma \sqrt{n}} \right)^p - \mathbf{E} (\sup_{t \geq 0} W(t))^p \right| \ll n^{-p/4} \log n.$$

*Remark.* In fact, Chassaing and Marckert [6] showed an even stronger result: In some probability space there exist a sequence of copies of  $w_n$  and a sequence of theta-distributed random variables  $D_n$  such that for any  $p \geq 1$

$$\left| \frac{2w_n}{\sigma \sqrt{n}} - D_n \right|_p = O\left(n^{-1/4} \sqrt{\log n}\right)$$

where the  $O$ -constant depends on  $p$ .

Recently, Chassaing, Marckert, and Yor [7] have used Theorems 1 and 4 in conjunction with results of Aldous [1] to obtain a weak limit theorem (without moments) for the joint law of height and width of simply generated trees. (For binary trees they present an elementary proof, too.)

## 2. PLAN OF THE PROOF OF THEOREM 3

In view of Theorem 2 the result of Theorem 3 is not unexpected. Nevertheless, it does not follow directly from Theorem 2 since convergence of moments is not automatically transferred via weak convergence (from Theorem 1).

In order to prove Theorem 3 we actually use the result of Theorem 1, that is, the normalized profile of Galton-Watson trees converges weakly to Brownian excursion local time:  $(l_n(t), t \geq 0) \xrightarrow{w} (l(t), t \geq 0)$ . However, we need some additional considerations: In [17] (see also [14]) Drmota and Marckert introduced the notion of so-called *polynomial convergence* (that is inspired by the notion of uniform integrability). The key property for our purposes is the following one. It generalizes the results of [17] (see also [14, Theorem 3.7]) that only apply for processes with compact support.

**Theorem 5.** *Let  $x_n(t)$  be a sequence of stochastic processes in  $C[0, \infty)$  which converges weakly to  $x(t)$ . Assume that for any choice of fixed positive integers  $p$  and  $d$  there exist positive constants  $c_0, c_1, c_2, c_3$  such that*

$$\sup_{n \geq 0} \mathbf{E} |x_n(t)|^p \leq c_0 e^{-c_1 t} \text{ for all } t \geq 0, \quad (5)$$

and

$$\sup_{n \geq 0} \mathbf{E} |x_n(t+s) - x_n(t)|^{2d} \leq c_2 e^{-c_3 t} s^d \text{ for all } s, t \geq 0. \quad (6)$$

Then  $x_n(t)$  is polynomially convergent to  $x(t)$ , that is, for every continuous functional  $F : C[0, \infty) \rightarrow \mathbb{R}$  of polynomial growth (i.e.  $|F(y)| \ll (1 + \|y\|_\infty)^r$  for some  $r \geq 0$ ) we have

$$\lim_{n \rightarrow \infty} \mathbf{E} F(x_n) = \mathbf{E} F(x).$$

We will show that  $l_n(t)$  satisfies the assumptions (5) and (6) of Theorem 5 and thus taking  $F(x) = \sup_{t \geq 0} x(t)^r$  yields immediately Theorem 3.

The next section is devoted to the proof of Theorem 5. In sections 4 and 5 we prove (5) and (6). Finally in section 6 we provide some further applications of Theorem 5.

## 3. PROOF OF THEOREM 5

Let us start with the following two observations.

**Lemma 1.** *Suppose that  $x_n(t)$  satisfies (5). Then for every  $p \geq 0$  we have*

$$\mathbf{E} \sup_{j \in \mathbb{N}} |x_n(j)|^p \ll 1$$

uniformly for all  $n$ .

*Proof.* Since  $\mathbf{E} |x_n(t)|^{p+1} \ll e^{-c_1 t}$ , uniformly in  $n$ , we have

$$\begin{aligned} \mathbf{P} \left\{ \sup_{j \in \mathbb{N}} |x_n(j)| \geq A \right\} &\leq \sum_{j \geq 0} \mathbf{P} \{ |x_n(j)| \geq A \} \\ &\leq \frac{1}{A^{p+1}} \sum_{j \geq 0} \mathbf{E} |x_n(j)|^{p+1} \text{ by Markov's inequality} \\ &\ll \frac{1}{A^{p+1}} \sum_{j \geq 0} e^{-c_1 j} \ll \frac{1}{A^{p+1}} \end{aligned}$$

Thus it follows that

$$\mathbf{E} \left( \sup_{j \in \mathbf{N}} |x_n(j)|^p \right) \ll 1 + p \int_1^\infty A^{p-1} \frac{1}{A^{p+1}} dA \ll 1.$$

□

**Lemma 2.** *Suppose that  $x_n(t)$  satisfies (6). Then, for fixed  $p$  we have*

$$\mathbf{E} \left( \sup_{|s-t| \leq \delta} |x_n(s) - x_n(t)|^p \right) \ll \delta^{p/2}.$$

uniformly for  $\delta$  with  $0 < \delta < 1$  and for all  $n$ .

*Proof.* First we prove that for every integer  $d > 1$  there exists a constant  $K > 0$  such that for  $\varepsilon > 0$  and  $0 < \delta < 1$

$$\mathbf{P} \left\{ \sup_{|s-t| \leq \delta} |x_n(s) - x_n(t)| \geq \varepsilon \right\} \leq K \frac{\delta^{d-1}}{\varepsilon^{2d}}. \quad (7)$$

Arguing as in [5, pp. 95] guarantees that there exists a constant  $K_1 > 0$  such that for all  $m \geq 0$

$$\mathbf{P} \left\{ \sup_{|s-t| \leq \delta, m \leq s, t \leq m+2} |x_n(s) - x_n(t)| \geq \varepsilon \right\} \leq K_1 e^{-c_5 m} \frac{\delta^{d-1}}{\varepsilon^{2d}}.$$

Thus

$$\mathbf{P} \left\{ \sup_{|s-t| \leq \delta} |x_n(s) - x_n(t)| \geq \varepsilon \right\} \leq \sum_{m=0}^{\infty} K_1 e^{-c_5 m} \frac{\delta^{d-1}}{\varepsilon^{2d}} \leq K \frac{\delta^{d-1}}{\varepsilon^{2d}}$$

for some constant  $K > 0$ .

Set

$$Z = \sup_{|s-t| \leq \delta} |x_n(s) - x_n(t)|.$$

Then by applying (7) it follows that (if  $2d \geq p + 1$ )

$$\begin{aligned} \mathbf{E} Z^p &= p \int_0^\infty z^{p-1} \mathbf{P}[Z > z] dz \\ &= p \int_0^{(K\delta)^{(d-1)/d}} z^{p-1} \mathbf{P}[Z > z] dz + p \int_{(K\delta)^{(d-1)/d}}^\infty z^{p-1} \mathbf{P}[Z > z] dz \\ &\leq (K\delta)^{p(d-1)/d} + pK\delta^{d-1} \int_{(K\delta)^{(d-1)/d}}^\infty z^{p-1-2d} dz \\ &\ll \delta^{p(d-1)/d} \leq \delta^{p/2}, \end{aligned}$$

which proves the Lemma. □

The proof of Theorem 5 is now an easy task. Note that the results of Lemma 1 and 2 in conjunction with the triangular inequality imply

$$\sup_{n \geq 0} \mathbf{E} \left( \sup_{t \geq 0} |x_n(t)|^r \right) < \infty \text{ for all } r \geq 0.$$

Thus, if  $F$  is a continuous functional of polynomial growth we have for any  $\varepsilon > 0$  we have

$$\sup_{n \geq 0} \mathbf{E} |F(x_n)|^{1+\varepsilon} < \infty.$$

By continuity of  $F$  we also obtain  $F(x_n) \xrightarrow{w} F(x)$  and finally, by Billingsley [4, p. 338] it directly follows that

$$\lim_{n \rightarrow \infty} \mathbf{E} F(x_n) = \mathbf{E} F(x)$$

as desired.  $\square$

#### 4. MOMENTS FOR THE PROFILE OF GALTON-WATSON TREES

We start with a lemma on the growth of coefficients of powers of certain generating functions.

**Lemma 3.** *Let  $z_0 \neq 0$  and  $\Delta = \{z : |z| < z_0 + \eta, |\arg(z - z_0)| > \vartheta\}$ , where  $\eta > 0$  and  $0 < \vartheta < \pi/2$ . Suppose that  $f(z)$  and  $g(z)$  are analytic functions in  $\Delta$  which satisfy*

$$|f(z)| \leq \exp\left(-C\sqrt{\left|1 - \frac{z}{z_0}\right|}\right), \quad z \in \Delta,$$

$$g(z) = 1 - D\sqrt{1 - \frac{z}{z_0}} + O\left(1 - \frac{z}{z_0}\right), \quad z \in \Delta,$$

for some positive constants  $C, D$ . Then for any fixed  $\ell$  there exists a constant  $C' > 0$  such that

$$[z^n] \frac{f(z)^r}{(1 - g(z))^\ell} = O\left(e^{-C'r/\sqrt{n}n^{(\ell-2)/2}}\right)$$

uniformly for all  $r, n \geq 0$  (where  $[z^n]F(z)$  denotes the coefficient of  $z^n$  of the function  $F(z)$ ).

*Proof.* The only difference to [23, Lemma 3.5] is the factor  $1/(1 - g(z))^\ell$ , but since its behavior in  $\Delta$  is known and [21, Theorem 3] is applicable, the proof is analogous to that of [23, Lemma 3.5].  $\square$

By means of this lemma we can show

**Lemma 4.** *For every fixed integer  $p > 0$  there exist positive constants  $c_0$  and  $c_1$  such that*

$$\sup_{n \geq 0} \mathbf{E} l_n(t)^p \leq c_0 e^{-c_1 t} \quad (8)$$

for all  $t \geq 0$ .

*Proof.* For technical simplicity we assume that  $g = \gcd\{i \geq 1 : \varphi_i > 0\} = 1$ . This assumption ensures that the tree function  $a(z)$  defined by (3) has only one singularity  $z_0 = 1/\varphi'(\tau)$  on the circle of convergence. If  $g = \gcd\{i \geq 0 : \varphi_i > 0\} > 1$  then we can use the substitution  $x = z^{1/g}$  to get  $a(z) = xb(x)$  where  $b(x)$  is analytic with only one singularity on the circle of convergence. Thus this case reduces to the case  $g = 1$ . The other possibility is to deal with the  $g$  singularities  $z_0 e^{2\pi i j/g}$ ,  $j = 0, 1, \dots, g-1$ , on the circle of convergence and add all contributions.

In particular, it is also well known that (if  $g = 1$ )  $a(z)$  admits a representation of the following kind

$$a(z) = \tau - \frac{\tau\sqrt{2}}{\sigma} \sqrt{1 - \frac{z}{z_0}} + O\left(\left|1 - \frac{z}{z_0}\right|\right), \quad (9)$$

that is valid for  $|z| < z_0 + \eta$  and  $\arg(z - z_0) \neq 0$ , where  $\eta > 0$  is suitably small, compare with [35] and [13].

In what follows we will need the local expansion of  $\alpha(z) = z\varphi'(a(z))$ . From (9) we immediately get

$$\alpha(z) = 1 - \sigma\sqrt{2}\sqrt{1 - \frac{z}{z_0}} + O\left(\left|1 - \frac{z}{z_0}\right|\right) \quad (10)$$

for  $|z| < z_0 + \eta$  and  $\arg(z - z_0) \neq 0$ .

Due to (10) there exists a constant  $C > 0$  such that  $|\alpha(z)| \leq \exp\left(-C\sqrt{|1 - z/z_0|}\right)$  for  $z \in \Delta$  (with  $\Delta$  from Lemma 3). Furthermore, it follows that

$$\sup_{z \in \Delta} |\alpha(z)| = 1, \quad (11)$$

where we have to choose  $\eta > 0$  and  $0 < \vartheta < \pi/2$  in a proper way. First, since the power series of  $\alpha(z)$  has only positive coefficients, we have  $\max_{|z| \leq z_0} |\alpha(z)| = 1$ . If we assume that  $d = \gcd\{i \geq 1 : \varphi_i > 0\} = 1$  it also follows that

$$\max_{|z| \leq z_0, |z - z_0| \geq \varepsilon} |\alpha(z)| < 1$$

for every  $\varepsilon > 0$ . Now, in the vicinity of the singularity  $z_0$ , that is, for  $|z - z_0| < \varepsilon$  we can again use (10) and get for  $z = z_0(1 + te^{i\theta})$

$$|\alpha(1 + te^{i\theta})| = \left|1 - \sigma\sqrt{2}te^{\pm i(\pi - \theta)/2} + O(t)\right|, \quad (12)$$

where  $\theta > \pi/2$ . Hence we have  $|\alpha(z)| \leq 1$  for  $|z - z_0| \leq \varepsilon$  and  $|\arg(z - z_0)| > \theta$ . Finally, for  $|z| \leq z_0 + \eta$  and  $|z - z_0| \geq \varepsilon$  we obtain the same inequality from (12) by a continuity argument (for some sufficiently small  $\eta > 0$ ). This proves (11).

Now observe that by substituting  $r = \lfloor t\sqrt{n} \rfloor$  (8) becomes

$$\mathbf{E} L_n(r)^p \leq c_0 e^{-c_1 r / \sqrt{n}} n^{p/2}. \quad (13)$$

Furthermore note that it suffices to show (13) for the  $p$ th factorial moment instead of the  $p$ th moment, which we can easily express in terms of the proper coefficient of a generating function. Indeed we have

$$\mathbf{E} [L_n(r)]_p = \frac{1}{a_n} [z^n] \left( \frac{\partial}{\partial u} \right)^p y_r(z, ua(z)) \Big|_{u=a(z)},$$

where

$$\begin{aligned} y_0(z, u) &= u \\ y_{i+1}(z, u) &= z\varphi(y_i(z, u)), \quad i \geq 0. \end{aligned} \quad (14)$$

In order to evaluate this coefficient we use Lemma 3 which translates the local behavior of the function near its singularity into an asymptotic estimate for the coefficients.

By [24, p. 287, equ. (22)] we have

$$\left( \frac{\partial}{\partial u} \right)^p y_r(z, ua(z)) \Big|_{u=a(z)} = O\left( a(z)^p |\alpha(z)|^r \left| \frac{1 - \alpha(z)^r}{1 - \alpha(z)} \right|^{p-1} \right) \quad (15)$$

From (11) we get

$$\max_{z \in \Delta} \left| \frac{1 - \alpha(z)^r}{1 - \alpha(z)} \right| \leq r. \quad (16)$$

Moreover  $a(z)^p$  behaves like a constant near the singularity and  $\alpha(z)^r$  meets the condition in Lemma 3. Hence the last factor in (15) is bounded by  $r^{p-1}$  and hence contributes a factor  $n^{(p-1)/2}$  to the order of magnitude of  $\mathbf{E} [L_n(r)]_p$ . Applying Lemma 3, which yields  $\exp(-c_1 r / \sqrt{n})$ , and normalizing by  $a_n \sim \tau / \sigma z_0^n \sqrt{2\pi n^3}$  we get the desired result.  $\square$

## 5. QUANTITATIVE TIGHTNESS ESTIMATES

With help of Lemma 3 we can prove the following quantitative *tightness estimate*.

**Lemma 5.** *For every fixed positive integer  $d$  there exist constants  $c_1, c_2$  such that for every  $s, t > 0$*

$$\mathbf{E} |l_n(t+s) - l_n(t)|^{2d} \leq c_1 e^{-c_2 t} s^d. \quad (17)$$

*Proof(Sketch).* Observe that (17) is equivalent to

$$\mathbf{E} |L_n(r) - L_n(r+h)|^{2d} \leq c_1 e^{-c_2 r / \sqrt{n}} h^d n^{d/2} \quad (18)$$

which is quite similar to [15, Theorem 6.1]. From [15] it follows that

$$\mathbf{E} |L_n(r) - L_n(r+h)|^{2d} = \frac{1}{a_n} [z^n] H_{rh}(z),$$

in which

$$H_{rh}(z) = \left( u \frac{\partial}{\partial u} \right)^{2d} y_r(z, u y_h(z, u^{-1} a(z))) \Big|_{u=1}$$

and  $y(z, u)$  is given by (14).

Evaluation of this coefficient is again done by Lemma 3. By [15, Proposition 6.1] it is easy to show that

$$H_{rh}(z) = \alpha(z)^r \sum_{j=0}^d G_{j,rh}(z) \frac{(1 - \alpha(z)^h)^j}{(1 - \alpha(z))^{d-1+j}}, \quad (19)$$

where  $G_{j,rh}(z)$  satisfy

$$\max_{z \in \Delta} |G_{j,rh}(z)| = O(1).$$

Eventually, an application of (16), with  $h$  instead of  $r$ , and Lemma 3 to (19) yields

$$[z^n] H_{rh}(z) = O\left(\frac{h^d n^{(d-3)/2}}{z_0^n}\right)$$

and, thus, by  $a_n \sim \tau / \sigma z_0^n \sqrt{2\pi n^3}$  the proof is complete.  $\square$

## 6. EXTENSIONS

**6.1. Nodes of given degree.** In [12] the number of nodes with fixed degree  $d$  in layers of random trees was investigated. In this case also limit theorems like Theorems 1 and 2 hold. In fact, we have

**Theorem 6.** *Let  $L_n^{(d)}(k)$  denote the number of nodes with degree  $d$  in layer  $k$  in a random tree of total progeny  $n$ . Furthermore, set*

$$l_n^{(d)}(t) = \frac{2}{\sigma c_d \sqrt{n}} L_n^{(d)}\left(\frac{2t}{\sigma} \sqrt{n}\right),$$

where  $c_d = \varphi_{d-1} \tau^{d-1} / \varphi(\tau)$ . Then we have

1.

$$l_n^{(d)}(t) \xrightarrow{w} l(t) \quad \text{and} \quad \sup_{t \geq 0} l_n^{(d)}(t) \xrightarrow{w} \sup_{t \geq 0} l(t)$$

2.

$$\mathbf{E} \left( \left( w_n^{(d)} \right)^p \right) = \mathbf{E} \left( \sup_{t \geq 0} l(t) \right)^p (1 + o(1)),$$

where  $w_n^{(d)} = \max_{k \geq 0} l_n^{(d)}(k)$ .

*Proof(Sketch).*

Part 1 was proved in [12]. The proof of part 2 runs similarly to the proof of Theorem 3. The only crucial point is to get estimates as in Lemma 4 and Lemma 5, namely

$$\mathbf{E} L_n^{(d)}(r)^p \leq c_1 e^{-c_2 r / \sqrt{n}} n^{p/2}$$

and

$$\mathbf{E} \left| L_n^{(d)}(r) - L_n^{(d)}(r+h) \right|^{2d} \leq c_1 e^{-c_2 r / \sqrt{n}} h^d n^{d/2}. \quad (20)$$

Both inequalities can be proved in a similar manner, so let us look at the second one (the first is the easier one). The results in [12] imply

$$\mathbf{E} \left| L_n^{(d)}(r) - L_n^{(d)}(r+h) \right|^{2d} = \frac{2}{\sigma a_n} [z^n] H_{2r/\sigma, 2h/\sigma}^{(d)}(z)$$

with

$$H_{rh}^{(d)}(z) = \left( u \frac{\partial}{\partial u} \right)^{2d} y_r(z, z(u-1)\varphi_{d-1} y_{h-1}(z, z(u^{-1}-1)\varphi_{d-1} a(z)^{d-1} + a(z)) \\ y_h(z, z(u^{-1}-1)\varphi_{d-1} a(z)^{d-1} + a(z)) \Big|_{u=1}$$

and since the right-hand side of this equation can be expressed in a form similar to (19), we can easily prove (20).  $\square$

**6.2. Strata of random mappings.** A random mapping of size  $n$  is an element of the set of all mappings of a set with  $n$  elements into itself equipped with the uniform distribution. These mappings can be represented by functional digraphs consisting of components which are cycles of trees. The set of points in distance  $r$  from a cycle is called the  $r$ th stratum of a random mapping. This parameter was previously studied in [2, 11, 16, 36, 39]. For general results on random mappings and literature see [32, 20]. Let  $M_n(r)$  denote the number of nodes in the  $r$ th stratum of a random mapping of size  $n$ . Then in [16] we proved

**Theorem 7.** Let  $B(t)$  denote reflecting Brownian bridge, i.e., a process on the interval  $[0, 1]$  which is identical in law to  $|\overline{W}(s) - s\overline{W}(1)|$  ( $\overline{W}(t)$  is the standard Brownian motion), and  $l^{(B)}(t)$  its local time, i.e.,

$$l^{(B)}(t) = \lim_{\varepsilon \rightarrow 0} \frac{1}{\varepsilon} \int_0^1 I_{[t, t+\varepsilon]}(B(s)) ds$$

Then we have

$$m_n(t) = \frac{2}{\sqrt{n}} M_n(2t\sqrt{n}) \xrightarrow{w} l^{(B)}(t)$$

in  $C[0, \infty)$ , as  $n \rightarrow \infty$ . Thus we also have

$$\sup_{t \geq 0} m_n(t) \xrightarrow{w} \sup_{t \geq 0} l^{(B)}(t).$$

By means of this we can show

**Theorem 8.** We have

$$\mathbf{E} \left( \left( \sup_{t \geq 0} m_n(t) \right)^p \right) = \mathbf{E} \left( \sup_{t \geq 0} l^{(B)}(t) \right)^p (1 + o(1)), \quad (21)$$

*Proof(Sketch).* Again the crucial point is to get proper estimates. From [16] it is an easy exercise to get

$$\mathbf{E} |M_n(r) - M_n(r+h)|^{2d} = \frac{2n!}{n^n} [z^n] H_{2r, 2h}(z),$$

in which

$$H_{rh}(z) = \left( u \frac{\partial}{\partial u} \right)^{2d} \frac{1}{1 - y_r(z, u y_h(z, u^{-1}a(z)))} \Big|_{u=1}.$$

This function can be written in a form similar to (19) and thus we can easily prove

$$\mathbf{E} |M_n(r) - M_n(r+h)|^{2d} \leq c_1 e^{-c_2 r / \sqrt{n}} h^d n^{d/2}$$

and then (21). The corresponding bound for the moments, obtained in the same way, carries out even easier.  $\square$

**6.3. Height of random trees.** The same method can be used to re-derive the analogue for the height  $h_n$  of simply generated trees (see Flajolet and Odlyzko [19]).

**Theorem 9.** Suppose that there exists a minimal positive solution  $\tau < R$  of  $t\varphi'(t) = \varphi(t)$ . Then

$$\mathbf{E} (h_n^p) = \left( \frac{\sqrt{2n}}{\sigma} \right)^p p(p-1) \Gamma\left(\frac{p}{2}\right) \zeta(p) (1 + o(1))$$

as  $n \rightarrow \infty$ .

$h_n$  is equal to the maximum of the traversal process  $T_n(r)$ , defined to be the distance between the root and the  $r$ th node during preorder traversal of the tree. Obviously, the same holds when we only traverse leaves (call the corresponding process  $\hat{T}_n(r)$ ). It is well known (see [1]) that

$$X_n(t) = \frac{1}{\sqrt{n}} T_n(2tn) \xrightarrow{w} \frac{2}{\sigma} W(t)$$

The height of leaves was investigated by several authors (see [30, 31, 26, 10, 23]). Here a similar limit theorem holds: With  $\hat{X}_n(t) = \hat{T}_n(tn)/\sqrt{n}$  we have (see [23])

$$\hat{X}_n\left(\frac{\varphi_0}{\varphi(\tau)}t\right) \xrightarrow{w} \frac{2}{\sigma}W(t).$$

In addition, in [23] the tightness estimate

$$P\{|\hat{X}_n(s) - \hat{X}_n(t)| \geq \varepsilon\} \leq C \frac{1}{\varepsilon^4|s-t|} \exp\left(-D \frac{\varepsilon}{\sqrt{|s-t|}}\right)$$

for some positive constants  $C$  and  $D$  was shown. This can be used to derive moment estimates like in Lemma 5 and then one proceeds as in the previous section to re-derive Flajolet and Odlyzko's [19] result on the moments of the height.

Finally, we want to mention that it is also possible to obtain the moments of the height of a random mapping (this was done by Flajolet and Odlyzko [20]) by our method. One has to use the weak limit theorem by Aldous and Pitman [2] and derive a tightness estimate in a similar fashion as has been done in [16].

#### REFERENCES

- [1] D. J. ALDOUS, The continuum random tree II: an overview, *Stochastic Analysis*, M. T. Barlow and N. H. Bingham, Eds., Cambridge University Press 1991, 23–70.
- [2] D. J. ALDOUS AND J. PITMAN, Brownian bridge asymptotics for random mappings, *Random Struct. Alg.* 5 (1994), 487–512.
- [3] P. BIANE AND M. YOR, Valeurs principales associées aux temps locaux Browniens, *Bull. Sci. Math.* 111 (1987), 23–101.
- [4] P. BILLINGSLEY, *Probability and Measure*, 3rd Ed., John Wiley & Sons, New York, 1995.
- [5] P. BILLINGSLEY, *Convergence of Probability Measures*, John Wiley & Sons, New York, 1968.
- [6] P. CHASSAING AND J.-F. MARCKERT, Parking functions, empirical processes, and the width of rooted labeled trees, *Electron. J. Combin.* 8, R14, 2001.
- [7] P. CHASSAING, J.-F. MARCKERT, AND M. YOR, The height and width of simple trees, in *Mathematics and computer science (Versailles, 2000)*, 17–30, Trends Math., Birkhäuser, Basel, 2000.
- [8] K. L. CHUNG, Excursions in Brownian motion, *Ark. Mat.* 14 (1976), 155–177.
- [9] J. W. COHEN AND G. HOOGHIEMSTRA, Brownian excursion, the  $M/M/1$  queue and their occupation times, *Mathematics of Operations Research* 6, 4 (1981), 608–629.
- [10] M. DRMOTA, The height distribution of leaves in rooted trees, *Discr. Math. Appl.* 4 (1994), 45–58 (translated from *Diskretn. Mat.* 6 (1994), 67–82).
- [11] M. DRMOTA, Correlations on the strata of a random mapping, *Random Struct. Alg.* 6 (1995), 357–365.
- [12] M. DRMOTA, On nodes of given degree in random trees, in *Probabilistic methods in discrete mathematics. Proceedings of the fourth international Petrozavodsk conference, Petrozavodsk, Russia, June 3–7, 1996*, Kolchin, V. F. (ed.) et al., Utrecht: VSP. 31–44, 1997.
- [13] M. DRMOTA, Systems of functional equations, *Random Struct. Alg.* 10 (1997), 103–124.
- [14] M. DRMOTA, Stochastic analysis of tree-like data structures, *Proc. R. Soc. Lond. A* 460 (2004), 271–307.
- [15] M. DRMOTA AND B. GITTENBERGER, On the profile of random trees, *Random Struct. Alg.* 10 (1997), 421–451.
- [16] M. DRMOTA AND B. GITTENBERGER, Strata of random mappings – a combinatorial approach, *Stoch. Proc. Appl.* 82 (1999), 157–171.
- [17] M. DRMOTA AND J.-F. MARCKERT, Reinforced weak convergence of stochastic processes, *Statistics Probab. Letters*, to appear.
- [18] R. T. DURRETT, D. L. IGLEHART AND D. R. MILLER, Weak convergence to Brownian meander and Brownian excursion, *Ann. Prob.* 5, (1977), 117–129.
- [19] P. FLAJOLET AND A. M. ODLYZKO, The average height of binary trees and other simple trees, *J. Comp. Syst. Sci.* 25 (1982), 171–213.

- [20] P. FLAJOLET AND A. M. ODLYZKO, Random mapping statistics, In *Advances in Cryptology* (1990), J.-J. Quisquater and J. Vandewalle, Eds., vol. 434 of *Lecture Notes in Computer Science*, Springer Verlag, pp. 329-354. Proceedings of EUROCRYPT '89, Houtalen, Belgium, April 1989.
- [21] P. FLAJOLET AND A. M. ODLYZKO, Singularity analysis of generating functions, *SIAM J. on Discrete Math.* 3, 2 (1990), 216–240.
- [22] R. K. GETTOOR AND M. J. SHARPE, Excursions of Brownian motion and Bessel processes, *Z. Wahrsch. verw. Gebiete* 47 (1979), 83–106.
- [23] B. GITTENBERGER, On the contour of random trees, *SIAM J. Discrete Math.* 12, No. 4 (1999), 434–458.
- [24] B. GITTENBERGER, On the profile of random forests, in *Mathematics and Computer Science II*, B. Chauvin, P. Flajolet, D. Gardy, and A. Mokkadem eds., Birkhuser, 2002, 279–293.
- [25] B. GITTENBERGER AND G. LOUCHARD, The Brownian excursion multidimensional local time density, *J. Appl. Probab.* 36 (2) (1999), 350-373.
- [26] W. GUTJAHN AND G. CH. PFLUG, The asymptotic contour process of a binary tree is a Brownian excursion, *Stochastic Processes and their Applications* 41 (1992), 69–89.
- [27] G. HOOGHIEMSTRA, On the explicit form of the density of Brownian excursion local time, *Proc. AMS* 84 (1982), 127–130.
- [28] D. P. KENNEDY, The distribution of the maximum Brownian excursion, *J. Appl. Prob.* 13 (1976), 371–376.
- [29] G. KERSTING, On the height profile of a conditioned Galton-Watson tree, 1998, unpublished.
- [30] P. KIRSCHENHOFER, On the height of leaves in binary trees, *J. Comb. Inf. Syst. Sci.* 8 (1983), 44–60.
- [31] P. KIRSCHENHOFER, Some new results on the average height of binary trees, *Ars Combinatoria* 16A (1983), 255–260.
- [32] V. F. KOLCHIN, *Random Mappings*, Optimization Software, New York, 1986.
- [33] J. KOMLOS, P. MAJOR, AND G. TUSNADY, An approximation of partial sums of independent RV's and the sample DF. II, *Z. Wahrscheinlichkeitstheor. Verw. Geb.* 34 (1976), 33–58.
- [34] G. LOUCHARD, Kac's formula, Levy's local time and Brownian excursion, *J. Appl. Prob.* 21 (1984), 479–499.
- [35] A. MEIR AND J. W. MOON, On the altitude of nodes in random trees, *Can. J. Math.* 30 (1978), 997–1015.
- [36] L. MUTAFCHIEV, The limit distribution of the number of nodes in low strata of a random mapping, *Stat. Prob. Lett.* 7 (1989), 247–251.
- [37] A. M. ODLYZKO AND H. WILF, Bandwidths and Profiles of Trees, *J. Combin. Theory, Ser. B* 42 (1987), 348–370.
- [38] J. PITMAN, The SDE solved by local times of a Brownian excursion or bridge derived from the height profile of a random tree, *Ann. Prob.* 27 (1999), 261–283.
- [39] G. V. PROSKURIN, On the distribution of the number of vertices in strata of a random mapping, *Theory Prob. Appl.* 18 (1973), 803–808.
- [40] L. TAKÁCS, Limit distributions for queues and random rooted trees, *J. Appl. Math. Stoch. Analysis* 6, No. 3, (1993), 189–216.
- [41] L. TAKÁCS, Brownian local times, *J. Appl. Math. Stoch. Analysis* 8, No. 3, (1995), 209–232.