A functional central limit theorem for branching random walks, with applications to Quicksort asymptotics

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> > based on joint work with

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Let  $X_n$  be the number of comparisons needed by QUICKSORT to sort the first n in a sequence  $(U_m)_{m \in \mathbb{N}}$  of independent uniforms, using the respective first element of the list as a pivot.

• Régnier (1989) 'found the martingale',

 $Y_n:=(X_n-EX_n)/(n+1)
ightarrow Y_\infty$  almost surely, as  $n
ightarrow\infty,$ 

- Rösler (1989) characterized  $\mathcal{L}(Y_{\infty})$  as a fixed point of a contraction,
- Neininger (2014) obtained an associated second order result,

 $\sqrt{n/(2\log n)}(Y_n - Y_\infty) \rightarrow N(0,1)$  in distribution,

using the contraction method. (Fuchs (2015): Method of moments) We will use the well-established link to branching random walks (Biggins, Chauvin, Devroye, Marckert, Roualt and others) to obtain

- yet another proof, indeed: of a stronger result,
- similar results for other (tree) models.

### **BRW: Basics**



Combines splitting (branching) and shifting (random walk). Basic parameter: (distribution of) a point process

$$\zeta = \sum_{i=1}^{N} \delta_{z_i}, \quad \text{where } \begin{cases} N & : \text{ (random) no. of children,} \\ z_1, \dots, z_N & : \text{ (random) location shifts,} \end{cases}$$

with i.i.d. copies of  $\zeta$  for the separate individuals.

- $(Z_n)_{n\in\mathbb{N}_0}$ : a BRW with parameter  $\mathcal{L}(\eta)$ ,
- ν with ν(A) := 𝔼ζ(A) is the intensity measure of ζ,
- $\nu_n(A) := \mathbb{E}Z_n(A)$  is the intensity measure of  $Z_n$ ,
- by construction,  $Z_1 = \eta$  and  $\nu_1 = \nu$ .

Let '^' denote the resp. moment generating function,

$$\hat{\zeta}(\theta) = \int e^{\theta \times} \zeta(dx), \quad \hat{Z}_n(\theta) = \int e^{\theta \times} Z_n(dx), \quad \hat{\nu}_n(\theta) = \int e^{\theta \times} \nu_n(dx).$$

Fundamental observation:

$$\nu_n = \nu^{\star n}, \ \hat{\nu}_n(\theta) = \nu(\theta)^n.$$

As a consequence,

$$W_n(\theta) := \hat{\nu}(\theta)^{-n} \hat{Z}_n(\theta), \ n \in \mathbb{N},$$

is a martingale.

(A) 
$$P(\zeta(\mathbb{R}) = 0) = 0$$
,  $P(\zeta(\mathbb{R}) < \infty) = 1$ ,  $P(\zeta(\mathbb{R}) = 1) < 1$ .  
(B) For some  $p_0 > 2$ ,  $\theta_0 > 0$ ,  
 $\mathbb{E}(\hat{\zeta}(\theta))^{p_0} < \infty$  for all  $\theta \in (-\theta_0, \theta_0)$ .

Let  $N_n := Z_n(\mathbb{R})$  be the (random) number of particles in generation n. (A) implies  $m := \mathbb{E}\zeta(\mathbb{R}) > 1$ , hence  $(N_n)_{n \in \mathbb{N}_0}$  is supercritical, so that

 $m^{-n}N_n \rightarrow N_\infty$  almost surely.

(A) and (B) also imply that  $W_n(\theta) \to W_\infty(\theta)$  almost surely, which we regard as a strong law of large numbers for the Biggins martingale. We will need

$$\sigma^2 := \operatorname{var}(N_\infty), \quad \tau^2 := (\log \hat{\nu})''(0).$$

Fix R > 0 and let  $\mathbb{D} := \{z \in \mathbb{C} : |z| \le R\}$ . Regard

$$D_n(u) := m^{n/2} \left( W_{\infty}\left(\frac{u}{\sqrt{n}}\right) - W_n\left(\frac{u}{\sqrt{n}}\right) \right)$$

as a random element of the space  $\mathbb{A}$  of continuous function  $f: \mathbb{D} \to \mathbb{C}$ that are analytic on  $\mathbb{D}^\circ$ . Endow  $\mathbb{A}$  with uniform convergence.

#### Theorem

As  $n \to \infty$ ,  $D_n$  converges in distribution to a limit process  $D_\infty$ . The distribution of  $D_\infty$  is the same as the distribution of

 $\mathbb{D} \ni u \mapsto \sigma \sqrt{N_{\infty}} \, \xi(\tau u),$ 

with  $N_{\infty}$ ,  $\xi$  independent and  $\xi$  the random analytic function

$$\xi(u) := \sum_{k=1}^{\infty} \xi_k \frac{u^k}{\sqrt{k!}}, \quad (\xi_k)_{k \in \mathbb{N}} \text{ i.i.d. } N(0,1).$$

- Switch to continuous time: particles have independent and exponentially distributed lifetimes.
- Generalize the FCLT from fixed times n = 1, 2, ... to stopping times τ<sub>1</sub>, τ<sub>2</sub>,... with τ<sub>n</sub> ↑ ∞.
- Use the continuous time BRW with  $\eta := 2 \, \delta_1$ .
- With τ<sub>n</sub> the time of birth of the *n*th particle, Z<sub>τn</sub> is the profile of the Quicksort tree of sublists.
- The punchline:
  - the quantity of interest is the mean of the profile,
  - on the transform side, mean means taking the derivative,
  - the functional  $\mathbb{A} \ni \phi \mapsto \phi'(0)$  is continuous.

Now use the FCLT and the Continuous Mapping Theorem.

- We may regard Neininger's result as a martingale CLT.
- It is known that these often hold in the stronger sense of stable convergence (Rényi).

Let  $(\mathcal{F}_n)_{n \in \mathbb{N}}$  be the martingale filtration. Basic idea:

# Instead of $\mathcal{L}(D_n)$ consider $\mathcal{L}[D_n|\mathcal{F}_n]$ .

- These conditional distributions are random variables with values in a space probability measures.
- The space of probability measures on the topological space A of analytic functions is itself a topological space  $\mathcal{M}$  if endowed with weak convergence.

## Theorem

In  $\mathcal{M}$ ,  $\mathcal{L}[D_n|\mathcal{F}_n]$  converges almost surely to  $\mathcal{L}(D_\infty)$ .

- Start with one black and one white ball. In each step, select one ball u.a.r., and add one ball of the same colour. Let X<sub>n</sub> be the number of white balls added after n steps, X = (X<sub>n</sub>)<sub>n∈N</sub>.
- $Y_n := X_n/(n+2)$ ,  $n \in \mathbb{N}$ , is a bd. martingale, hence  $Y_n \to Y_\infty$  a.s..
- Boundary theory (or direct calculation) says that X, given Y<sub>∞</sub> = θ, is a simple random walk on Z, with θ the probability for a move to the right.

• Let 
$$W_n := \sqrt{n}(Y_{\infty} - Y_n)$$
. Standard CLT gives

 $\mathcal{L}[W_n|Y_{\infty} = \theta] \rightarrow N(0, \theta(1-\theta))$  in distribution.

• We may rewrite this as a.s. convergence of the random p.m.  $\mathcal{L}[W_n|Y_\infty]$  to the random (!) p.m.  $N(0, Y_\infty(1 - Y_\infty))$ .

#### Theorem

 $\mathcal{L}[W_n|X_1,\ldots,X_n] \text{ converges a.s. to } N(0,Y_\infty(1-Y_\infty)).$  (Proof uses some beta-gamma algebra.)