# A functional central limit theorem for branching random walks, with applications to Quicksort asymptotics 

Rudolf Grübel<br>Leibniz Universität Hannover<br>based on joint work with<br>Zakhar Kabluchko<br>Westfälische Wilhelms Universität Münster

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## Background and history

Let $X_{n}$ be the number of comparisons needed by Quicksort to sort the first $n$ in a sequence $\left(U_{m}\right)_{m \in \mathbb{N}}$ of independent uniforms, using the respective first element of the list as a pivot.

- Régnier (1989) 'found the martingale',

$$
Y_{n}:=\left(X_{n}-E X_{n}\right) /(n+1) \rightarrow Y_{\infty} \text { almost surely, as } n \rightarrow \infty,
$$

- Rösler (1989) characterized $\mathcal{L}\left(Y_{\infty}\right)$ as a fixed point of a contraction,
- Neininger (2014) obtained an associated second order result,

$$
\sqrt{n /(2 \log n)}\left(Y_{n}-Y_{\infty}\right) \rightarrow N(0,1) \text { in distribution, }
$$

using the contraction method. (Fuchs (2015): Method of moments)
We will use the well-established link to branching random walks (Biggins, Chauvin, Devroye, Marckert, Roualt and others) to obtain

- yet another proof, indeed: of a stronger result,
- similar results for other (tree) models.


## BRW: Basics



Combines splitting (branching) and shifting (random walk).
Basic parameter: (distribution of) a point process

$$
\zeta=\sum_{i=1}^{N} \delta_{z_{j}}, \quad \text { where } \begin{cases}N & :(\text { random) no. of children, } \\ z_{1}, \ldots, z_{N} & :(\text { random) location shifts, }\end{cases}
$$

with i.i.d. copies of $\zeta$ for the separate individuals.

## BRW: The Biggins martingale

- $\left(Z_{n}\right)_{n \in \mathbb{N}_{0}}$ : a BRW with parameter $\mathcal{L}(\eta)$,
- $\nu$ with $\nu(A):=\mathbb{E} \zeta(A)$ is the intensity measure of $\zeta$,
- $\nu_{n}(A):=\mathbb{E} Z_{n}(A)$ is the intensity measure of $Z_{n}$,
- by construction, $Z_{1}=\eta$ and $\nu_{1}=\nu$.

Let "^’ denote the resp. moment generating function,

$$
\hat{\zeta}(\theta)=\int e^{\theta x} \zeta(d x), \quad \hat{Z}_{n}(\theta)=\int e^{\theta x} Z_{n}(d x), \quad \hat{\nu}_{n}(\theta)=\int e^{\theta x} \nu_{n}(d x) .
$$

Fundamental observation:

$$
\nu_{n}=\nu^{\star n}, \hat{\nu}_{n}(\theta)=\nu(\theta)^{n}
$$

As a consequence,

$$
W_{n}(\theta):=\hat{\nu}(\theta)^{-n} \hat{Z}_{n}(\theta), n \in \mathbb{N}
$$

is a martingale.

## BRW: Assumptions

(A) $P(\zeta(\mathbb{R})=0)=0, P(\zeta(\mathbb{R})<\infty)=1, P(\zeta(\mathbb{R})=1)<1$.
(B) For some $p_{0}>2, \theta_{0}>0$,

$$
\mathbb{E}(\hat{\zeta}(\theta))^{p_{0}}<\infty \text { for all } \theta \in\left(-\theta_{0}, \theta_{0}\right)
$$

Let $N_{n}:=Z_{n}(\mathbb{R})$ be the (random) number of particles in generation $n$. (A) implies $m:=\mathbb{E} \zeta(\mathbb{R})>1$, hence $\left(N_{n}\right)_{n \in \mathbb{N}_{0}}$ is supercritical, so that

$$
m^{-n} N_{n} \rightarrow N_{\infty} \text { almost surely. }
$$

(A) and (B) also imply that $W_{n}(\theta) \rightarrow W_{\infty}(\theta)$ almost surely, which we regard as a strong law of large numbers for the Biggins martingale.
We will need

$$
\sigma^{2}:=\operatorname{var}\left(N_{\infty}\right), \quad \tau^{2}:=(\log \hat{\nu})^{\prime \prime}(0)
$$

## BRW: A functional central limit theorem

Fix $R>0$ and let $\mathbb{D}:=\{z \in \mathbb{C}:|z| \leq R\}$. Regard

$$
D_{n}(u):=m^{n / 2}\left(W_{\infty}\left(\frac{u}{\sqrt{n}}\right)-W_{n}\left(\frac{u}{\sqrt{n}}\right)\right)
$$

as a random element of the space $\mathbb{A}$ of continuous function $f: \mathbb{D} \rightarrow \mathbb{C}$ that are analytic on $\mathbb{D}^{\circ}$. Endow $\mathbb{A}$ with uniform convergence.
Theorem
As $n \rightarrow \infty, D_{n}$ converges in distribution to a limit process $D_{\infty}$. The distribution of $D_{\infty}$ is the same as the distribution of

$$
\mathbb{D} \ni u \mapsto \sigma \sqrt{N_{\infty}} \xi(\tau u),
$$

with $N_{\infty}, \xi$ independent and $\xi$ the random analytic function

$$
\xi(u):=\sum_{k=1}^{\infty} \xi_{k} \frac{u^{k}}{\sqrt{k!}}, \quad\left(\xi_{k}\right)_{k \in \mathbb{N}} \text { i.i.d. } N(0,1) .
$$

## From BRW to Quicksort

- Switch to continuous time: particles have independent and exponentially distributed lifetimes.
- Generalize the FCLT from fixed times $n=1,2, \ldots$ to stopping times $\tau_{1}, \tau_{2}, \ldots$ with $\tau_{n} \uparrow \infty$.
- Use the continuous time BRW with $\eta:=2 \delta_{1}$.
- With $\tau_{n}$ the time of birth of the $n$th particle, $Z_{\tau_{n}}$ is the profile of the Quicksort tree of sublists.
- The punchline:
- the quantity of interest is the mean of the profile,
- on the transform side, mean means taking the derivative,
- the functional $\mathbb{A} \ni \phi \mapsto \phi^{\prime}(0)$ is continuous.

Now use the FCLT and the Continuous Mapping Theorem.

## The strengthening

- We may regard Neininger's result as a martingale CLT.
- It is known that these often hold in the stronger sense of stable convergence (Rényi).
Let $\left(\mathcal{F}_{n}\right)_{n \in \mathbb{N}}$ be the martingale filtration.
Basic idea:

$$
\text { Instead of } \mathcal{L}\left(D_{n}\right) \text { consider } \mathcal{L}\left[D_{n} \mid \mathcal{F}_{n}\right] \text {. }
$$

- These conditional distributions are random variables with values in a space probability measures.
- The space of probability measures on the topological space $\mathbb{A}$ of analytic functions is itself a topological space $\mathcal{M}$ if endowed with weak convergence.

Theorem
In $\mathcal{M}, \mathcal{L}\left[D_{n} \mid \mathcal{F}_{n}\right]$ converges almost surely to $\mathcal{L}\left(D_{\infty}\right)$.

## Back to basics: the Pólya urn

- Start with one black and one white ball. In each step, select one ball u.a.r., and add one ball of the same colour. Let $X_{n}$ be the number of white balls added after $n$ steps, $X=\left(X_{n}\right)_{n \in \mathbb{N}}$.
- $Y_{n}:=X_{n} /(n+2), n \in \mathbb{N}$, is a bd. martingale, hence $Y_{n} \rightarrow Y_{\infty}$ a.s..
- Boundary theory (or direct calculation) says that $X$, given $Y_{\infty}=\theta$, is a simple random walk on $\mathbb{Z}$, with $\theta$ the probability for a move to the right.
- Let $W_{n}:=\sqrt{n}\left(Y_{\infty}-Y_{n}\right)$. Standard CLT gives

$$
\mathcal{L}\left[W_{n} \mid Y_{\infty}=\theta\right] \rightarrow N(0, \theta(1-\theta)) \text { in distribution. }
$$

- We may rewrite this as a.s. convergence of the random p.m. $\mathcal{L}\left[W_{n} \mid Y_{\infty}\right]$ to the random (!) p.m. $N\left(0, Y_{\infty}\left(1-Y_{\infty}\right)\right)$.


## Theorem

$\mathcal{L}\left[W_{n} \mid X_{1}, \ldots, X_{n}\right]$ converges a.s. to $N\left(0, Y_{\infty}\left(1-Y_{\infty}\right)\right)$.
(Proof uses some beta-gamma algebra.)

